

Parameter estimation for subcritical affine processes

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An affine two-factor model

For $a > 0$, $b, \theta, m \in \mathbb{R}$ and $\alpha \in (1, 2]$, let us consider the SDE

$$\begin{cases} dY_t = (a - bY_t) dt + \sqrt[{\alpha}]{Y_t} dL_t, \\ dX_t = (m - \vartheta X_t) dt + \sqrt{Y_t} dB_t, \end{cases} \quad t \geq 0,$$

where $(L_t)_{t \geq 0}$ is a spectrally positive α -stable Lévy process with Lévy measure $C_\alpha z^{-1-\alpha} \mathbf{1}_{\{z > 0\}}$ with $C_\alpha := (\alpha \Gamma(-\alpha))^{-1}$ in case $\alpha \in (1, 2)$, a Brownian motion in case $\alpha = 2$, and $(B_t)_{t \geq 0}$ is an independent Brownian motion.

Existence and uniqueness of a strong solution

Let (Y_0, X_0) be \mathcal{F}_0 -measurable with $\mathbb{P}(Y_0 \geq 0) = 1$. Then there is a unique strong solution $(Y_t, X_t)_{t \geq 0}$ with $\mathbb{P}(Y_t \geq 0 \text{ for all } t \geq 0) = 1$. Moreover, $(Y_t, X_t)_{t \geq 0}$ is an affine process with infinitesimal generator

$$(\mathcal{A}f)(y, x) = (a - by)f'_1(y, x) + (m - \vartheta x)f'_2(y, x) + \frac{1}{2}yf''_{2,2}(y, x) + y \int_0^\infty (f(y+z, x) - f(y, x) - zf'_1(y, x)) \frac{C_\alpha}{z^{1+\alpha}} dz$$

in case of $\alpha \in (1, 2)$, and

$$(\mathcal{A}f)(y, x) = (a - by)f'_1(y, x) + (m - \vartheta x)f'_2(y, x) + \frac{1}{2}y(f''_{1,1}(y, x) + f''_{2,2}(y, x))$$

in case of $\alpha = 2$, where $f \in \mathcal{C}_c^2(\mathbb{R}_+ \times \mathbb{R}, \mathbb{R})$.

Stationarity in the subcritical case

Let $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 \geq 0) = 1$.

1 Then $(Y_t, X_t) \xrightarrow{\mathcal{L}} (Y_\infty, X_\infty)$ as $t \rightarrow \infty$, and the distribution of (Y_∞, X_∞) is given by

$$\mathbb{E}(e^{-\lambda_1 Y_\infty + i\lambda_2 X_\infty}) = \exp \left\{ -a \int_0^\infty v_s(\lambda_1, \lambda_2) ds + i \frac{m}{\theta} \lambda_2 \right\}$$

for $(\lambda_1, \lambda_2) \in \mathbb{R}_+ \times \mathbb{R}$, where $v_t(\lambda_1, \lambda_2)$, $t \geq 0$, is the unique non-negative solution of

$$\begin{cases} \frac{\partial v_t}{\partial t}(\lambda_1, \lambda_2) = -bv_t(\lambda_1, \lambda_2) - \frac{1}{\alpha}(v_t(\lambda_1, \lambda_2))^\alpha + \frac{1}{2}e^{-2\vartheta t} \lambda_2^2, \\ v_0(\lambda_1, \lambda_2) = \lambda_1. \end{cases} \quad t \geq 0,$$

2 If $(Y_0, X_0) \stackrel{\mathcal{L}}{=} (Y_\infty, X_\infty)$, then $(Y_t, X_t)_{t \geq 0}$ is strictly stationary.

3 If $\alpha = 2$, then (Y_∞, X_∞) is absolutely continuous, and

$$\mathbb{E}(e^{-\lambda_1 Y_\infty}) = \left(1 + \frac{\lambda_1}{2b}\right)^{-2a}, \quad \lambda_1 \in \mathbb{R}_+.$$

Ergodicity in the subcritical case if $\alpha = 2$

If $\alpha = 2$, $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 \geq 0) = 1$ then

$$\mathbb{P} \left(\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T f(Y_s, X_s) ds = \mathbb{E} f(Y_\infty, X_\infty) \right) = 1$$

for all Borel measurable functions $f: \mathbb{R}^2 \rightarrow \mathbb{R}$ with $\mathbb{E}|f(Y_\infty, X_\infty)| < \infty$.

It suffices to check a Foster–Lyapunov criteria :

- $(Y_t, X_t)_{t \geq 0}$ is a right process;
- all compact sets are petite for some skeleton chain;
- there exist $c > 0$ and $d \in \mathbb{R}$ such that

$$(\mathcal{A}_n V)(y, x) \leq -cV(y, x) + d, \quad (y, x) \in O_n, \quad n \in \mathbb{N},$$

where $O_n := \{(y, x) \in \mathbb{R}_+ \times \mathbb{R} : \|(y, x)\| < n\}$ for each $n \in \mathbb{N}$,

$$V(y, x) := (y - c_1)^2 + (x - c_2)^2, \quad (y, x) \in \mathbb{R}_+ \times \mathbb{R},$$

with some appropriate $c_1, c_2 \in \mathbb{R}$, and \mathcal{A}_n denotes the extended generator of the stopped process $(Y_t^{(n)}, X_t^{(n)})_{t \geq 0}$ given by

$$(Y_t^{(n)}, X_t^{(n)}) := \begin{cases} (Y_t, X_t), & \text{for } t < \tau_n, \\ (0, n), & \text{for } t \geq \tau_n, \end{cases}$$

where the stopping time τ_n is defined by

$$\tau_n := \inf\{t \in \mathbb{R}_+ : (Y_t, X_t) \in (\mathbb{R}_+ \times \mathbb{R}) \setminus O_n\}.$$

Existence and uniqueness of MLE of (ϑ, m) if $\alpha = 2$

If $\alpha = 2$, $a \geq \frac{1}{2}$, $\mathbb{P}(Y_0 > 0) = 1$ and $T > 0$, then there exists a unique MLE of (ϑ, m) based on observations $(Y_t, X_t)_{t \in [0, T]}$, and it is given by

$$\begin{bmatrix} \hat{\vartheta}_T \\ \hat{m}_T \end{bmatrix} = \begin{bmatrix} \int_0^T \frac{X_s^2}{Y_s} ds & -\int_0^T \frac{X_s}{Y_s} ds \\ -\int_0^T \frac{X_s}{Y_s} ds & \int_0^T \frac{1}{Y_s} ds \end{bmatrix}^{-1} \begin{bmatrix} -\int_0^T \frac{X_s}{Y_s} dX_s \\ \int_0^T \frac{1}{Y_s} dX_s \end{bmatrix}, \quad T > 0,$$

whenever $\int_0^T \frac{X_s^2}{Y_s} ds \int_0^T \frac{1}{Y_s} ds - \left(\int_0^T \frac{X_s}{Y_s} ds\right)^2 > 0$.

Consistency of MLE in the subcritical case if $\alpha = 2$

If $\alpha = 2$, $a \geq \frac{1}{2}$, $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 > 0) = 1$, then

$$\mathbb{P} \left(\lim_{T \rightarrow \infty} (\hat{\vartheta}_T, \hat{m}_T) = (\vartheta, m) \right) = 1.$$

Asymptotics of MLE in the subcritical case if $\alpha = 2$

If $\alpha = 2$, $a \geq \frac{1}{2}$, $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 > 0) = 1$, then

$$\begin{bmatrix} \sqrt{T}(\hat{\vartheta}_T - \vartheta) \\ \sqrt{T}(\hat{m}_T - m) \end{bmatrix} \xrightarrow{\mathcal{L}} \mathcal{N}_2(0, \Sigma_{\text{MLE}}) \quad \text{as } T \rightarrow \infty,$$

with

$$\Sigma_{\text{MLE}} := \frac{1}{\mathbb{E} \left(\frac{1}{Y_\infty} \right) \mathbb{E} \left(\frac{X_\infty^2}{Y_\infty} \right) - \left(\mathbb{E} \left(\frac{X_\infty}{Y_\infty} \right) \right)^2} \begin{bmatrix} \mathbb{E} \left(\frac{1}{Y_\infty} \right) & \mathbb{E} \left(\frac{X_\infty}{Y_\infty} \right) \\ \mathbb{E} \left(\frac{X_\infty}{Y_\infty} \right) & \mathbb{E} \left(\frac{X_\infty^2}{Y_\infty} \right) \end{bmatrix}.$$

LSE of (ϑ, m) based on discrete observations

LSE of (ϑ, m) based on observations X_0, X_1, \dots, X_n , $n \in \mathbb{N}$, is defined as

$$\arg \min_{(\vartheta, m) \in \mathbb{R}^2} \sum_{i=1}^n (X_i - X_{i-1} - (m - \vartheta X_{i-1}))^2, \quad n \in \mathbb{N},$$

and it is given by

$$\begin{bmatrix} \hat{\vartheta}_T \\ \hat{m}_T \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^n X_{i-1}^2 & -\sum_{i=1}^n X_{i-1} \\ -\sum_{i=1}^n X_{i-1} & n \end{bmatrix}^{-1} \begin{bmatrix} -\sum_{i=1}^n (X_i - X_{i-1})X_{i-1} \\ X_n - X_0 \end{bmatrix}, \quad n \in \mathbb{N},$$

whenever $n \sum_{i=1}^n X_{i-1}^2 - \left(\sum_{i=1}^n X_{i-1}\right)^2 > 0$.

LSE of (ϑ, m) based on continuous observations

LSE of (ϑ, m) based on observations $(X_t)_{t \in [0, T]}$, $T > 0$, is defined by

$$\begin{bmatrix} \hat{\vartheta}_T \\ \hat{m}_T \end{bmatrix} = \begin{bmatrix} \int_0^T \frac{X_s^2}{Y_s} ds & -\int_0^T \frac{X_s}{Y_s} ds \\ -\int_0^T \frac{X_s}{Y_s} ds & T \end{bmatrix}^{-1} \begin{bmatrix} -\int_0^T \frac{X_s}{Y_s} dX_s \\ X_T - X_0 \end{bmatrix}, \quad T > 0,$$

whenever $T \int_0^T \frac{X_s^2}{Y_s} ds - \left(\int_0^T \frac{X_s}{Y_s} ds\right)^2 > 0$, which holds almost surely.

Consistency of LSE in the subcritical case if $\alpha = 2$

If $\alpha = 2$, $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 > 0) = 1$, then

$$\mathbb{P} \left(\lim_{T \rightarrow \infty} (\hat{\vartheta}_T, \hat{m}_T) = (\vartheta, m) \right) = 1.$$

Asymptotics of LSE in the subcritical case if $\alpha = 2$

If $\alpha = 2$, $b > 0$, $\vartheta > 0$ and $\mathbb{P}(Y_0 > 0) = 1$, then

$$\begin{bmatrix} \sqrt{T}(\hat{\vartheta}_T - \vartheta) \\ \sqrt{T}(\hat{m}_T - m) \end{bmatrix} \xrightarrow{\mathcal{L}} \mathcal{N}_2(0, \Sigma_{\text{LSE}}) \quad \text{as } T \rightarrow \infty,$$

with some explicitly given Σ_{LSE} .

References

- [1] M. BARCZY, L. DÖRING, Z. LI and G. PAP (2013). Ergodicity for an affine two factor model. ArXiv: <http://arxiv.org/abs/1302.2534>
- [2] M. BARCZY, L. DÖRING, Z. LI and G. PAP (2013). Parameter estimation for an affine two factor model. ArXiv: <http://arxiv.org/abs/1302.3451>